Neural Network Determination of Parton Distribution Functions

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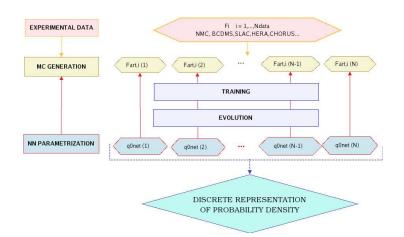
Issue and Standard Approach

- Given a set of data points we must determine a set of functions with error.
- We need an error band in the space of functions, i.e. a probability density $\mathcal{P}[q(x)]$ in the space of PDFs, q(x). For an observable \mathcal{F} depending on PDFs:

$$\langle \mathcal{F}[q(x)] \rangle = \int [\mathcal{D}q] \, \mathcal{F}[q(x)] \mathcal{P}[q(x)]$$

- Standard approach, choose a basis of functions and project PDFs on it: the ∞-dimensional space of function reduces to a finite-dimensional space of parameters.
- Issues:
 - Non trivial propagation of errors: non-gaussian errors and incompatible data.
 - The error associated to the choice of parametrisation is difficult to assess.

NNPDF approach



$$\langle \mathcal{F}[q(x)] \rangle = \int [\mathcal{D}q] \, \mathcal{F}[q(x)] \mathcal{P}[q(x)] \longrightarrow \langle \mathcal{F}[q(x)] \rangle = \frac{1}{N_{\mathrm{rep}}} \sum_{k=1}^{N_{\mathrm{rep}}} \mathcal{F}[q^{(k)(\mathrm{net})}(x)]$$

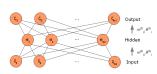
Main Ingredients

Monte Carlo determination of errors:

After fitting, the error of an observable depending on PDFs \rightarrow

$$\sigma_{\mathcal{F}[q(x)]} = \sqrt{\langle \mathcal{F}[q(x)]^2 \rangle - \langle \mathcal{F}[q(x)] \rangle^2}$$

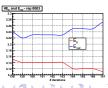
• Neural Networks as redundant and unbiased parametrisation of PDFs:



- * Each neuron receives input from neurons in preceding layer.
- * Activation determined by weights and thresholds according to a non linear function:

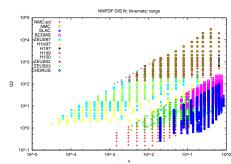
$$\xi_i = g(\sum_j \omega_{ij}\xi_j - \theta_i), \qquad g(x) = \frac{1}{1 + e^{-x}}$$

- Dynamical stopping criterion in order to fit data and not statistical noise.
- * Divide data in two sets: training and validation.
- * Minimisation is performed only on the training set. The validation χ^2 for the set is computed.
- * When the training χ^2 still decreases while the validation χ^2 stops decreasing \rightarrow STOP.



Singlet fit

- NLO fit.
- ZM-VFN treatment of heavy quarks.
- All DIS data included.
- Flavor Assumptions:
 - Symmetric strange sea $s(x) = \overline{s}(x)$
 - Strange sea proportional to non-strange sea $\bar{s}(x) = \frac{C}{2}(\bar{u}(x) + \bar{d}(x))$ (C = 0.5)

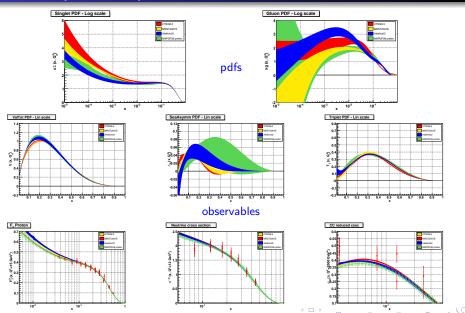


• Parametrization of 4+1 combinations of PDFs at $Q_0^2 = 2 \text{ GeV}^2$:

Singlet: $\Sigma(x)$ $\longmapsto NN_{\Sigma}(x)$ 2-3-2-1 20 pars Gluon: g(x) $\longmapsto NN_g(x)$ 2-3-2-1 20 pars Total valence: $V(x) \equiv u_V(x) + d_V(x) \longmapsto NN_V(x)$ 2-3-2-1 20 pars Non-singlet triplet: $T_3(x)$ $\longmapsto NN_{T3}(x)$ 2-3-2-1 20 pars Sea asymmetry: $\Delta_5(x) \equiv \bar{d}(x) - \bar{u}(x) \longmapsto NN_{\Delta}(x)$ 2-3-1 13 pars

93 parameters

Some Very Preliminary Results



Conclusions

- Standard approaches to PDFs fitting might lead to underestimation of errors associated with parton densities.
- Combination of Monte Carlo techniques and Neural Networks as unbiased interpolating functions has proved to be a fast and robust alternative method
- A non singlet fit has been published [hep-ph/0701127] and a full DIS fit will be published very soon.

BACKUP SLIDES

MC replicas of experimental data

• Generate a N_{rep} Monte Carlo sets of artificial data, or "pseudo-data" of the original N_{data} data points

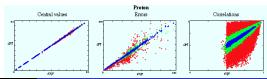
$$F_i^{(art)(k)}$$
 $i = 1, ..., N_{data}$ $k = 1, ..., N_{rep}$

according to:

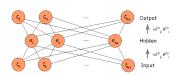
$$F_i^{(art)(k)} = (1 + r_N^{(k)})[F_i^{\text{exp}} + r_S^{(K)}\sigma_i^{\text{stat}} + \sum_{j=1}^{N_{ ext{sys}}} r_{j, ext{SY}}^{(k)}\sigma_{ij}^{ ext{sys}}]$$

- σ_i , experimental errors;
- r_i , zero mean gaussian random numbers distributed according to the experimental correlation matrix.
- ullet Validate MC replicas according to experimental data (statistical estimators, faithful representation of errors, convergence rate increasing $N_{\rm rep}$).

How many replicas do we need? 1000 replicas are enough to reproduce correlation to percent accuracy.



Explicit functional form of a NN



- Each neuron receives input from neurons in preceding layer.
- Activation determined by weight and threshold according to a non linear function:

$$\xi_i = g(\sum_j \omega_{ij}\xi_j - \theta_i), \qquad g(x) = \frac{1}{1 + e^{-x}}$$

• In a simple case (1-2-1) we have,

$$\xi_1^{(3)} = \frac{1}{\theta_1^{(3)} - \frac{\omega_{11}^{(2)}}{\frac{\omega_{11}^{(2)} - \xi_1^{(1)}\omega_{11}^{(1)}}{1 + e^{\theta_1^{(2)} - \xi_1^{(1)}\omega_{11}^{(1)}} - \frac{\omega_{12}^{(2)}}{\frac{\omega_{12}^{(2)} - \xi_1^{(1)}\omega_{21}^{(1)}}{1 + e^{\theta_2^{(2)} - \xi_1^{(1)}\omega_{21}^{(1)}}}}$$

- NNs are just another set of basis functions.
- Thanks to non linear behaviour, any function can be represented by a sufficiently big NN.



Statistical Estimators I: observables

Central value of the i-th experimental point

$$\left\langle F_i^{(\mathrm{art})} \right\rangle_{\mathrm{rep}} = \frac{1}{N_{\mathrm{rep}}} \sum_{k=1}^{N_{\mathrm{rep}}} F_i^{(\mathrm{art})(k)} \; .$$

Variance of the i-th experimental point

$$\sigma_i^{(\mathrm{art})} = \sqrt{\left\langle \left(F_i^{(\mathrm{art})} \right)^2 \right\rangle_{\mathrm{rep}} - \left\langle F_i^{(\mathrm{art})} \right\rangle_{\mathrm{rep}}^2} \; .$$

Associated covariance:

$$\begin{split} \rho_{ij}^{(\mathrm{art})} &= \frac{\left\langle F_i^{(\mathrm{art})} F_j^{(\mathrm{art})} \right\rangle_{\mathrm{rep}} - \left\langle F_i^{(\mathrm{art})} \right\rangle_{\mathrm{rep}} \left\langle F_j^{(\mathrm{art})} \right\rangle_{\mathrm{rep}}}{\sigma_i^{(\mathrm{art})} \sigma_j^{(\mathrm{art})}} \; . \\ &\quad \mathrm{cov}_{ij}^{(\mathrm{art})} = \rho_{ij}^{(\mathrm{art})} \sigma_i^{(\mathrm{art})} \sigma_j^{(\mathrm{art})}. \end{split}$$

Statistical Estimators II: replicas vs data

ullet Mean variance and percentage error on central values over the $N_{
m dat}$ data points.

$$\left\langle V \left[\left\langle F^{(\mathrm{art})} \right\rangle_{\mathrm{rep}} \right] \right\rangle_{\mathrm{dat}} = \frac{1}{N_{\mathrm{dat}}} \sum_{i=1}^{N_{\mathrm{dat}}} \left(\left\langle F_i^{(\mathrm{art})} \right\rangle_{\mathrm{rep}} - F_i^{(\mathrm{exp})} \right)^2 \; ,$$

$$\left\langle PE\left[\left\langle F^{(\mathrm{art})}\right\rangle_{\mathrm{rep}}\right]
ight
angle_{\mathrm{dat}} = rac{1}{N_{\mathrm{dat}}}\sum_{i=1}^{N_{\mathrm{dat}}}\left[rac{\left\langle F_{i}^{(\mathrm{art})}\right\rangle_{\mathrm{rep}} - F_{i}^{(\mathrm{exp})}}{F_{i}^{(\mathrm{exp})}}
ight] \;.$$

$$\bullet \ \left\langle V \left[\left\langle \sigma^{(\mathrm{art})} \right\rangle_{\mathrm{rep}} \right] \right\rangle_{\mathrm{dat}}, \left\langle V \left[\left\langle \rho^{(\mathrm{art})} \right\rangle_{\mathrm{rep}} \right] \right\rangle_{\mathrm{dat}}, \left\langle V \left[\left\langle \mathrm{cov}^{(\mathrm{art})} \right\rangle_{\mathrm{rep}} \right] \right\rangle_{\mathrm{dat}}$$

$$\left\langle \textit{PE}\left[\left\langle \sigma^{(\mathrm{art})}\right\rangle_{\mathrm{rep}}\right]\right\rangle_{\mathrm{dat}}, \left\langle \textit{PE}\left[\left\langle \rho^{(\mathrm{art})}\right\rangle_{\mathrm{rep}}\right]\right\rangle_{\mathrm{dat}}, \left\langle \textit{PE}\left[\left\langle \mathrm{cov^{(art)}}\right\rangle_{\mathrm{rep}}\right]\right\rangle_{\mathrm{dat}}$$

relative to errors, correlations and covariances are defined in the same way.

 These estimators indicate how close are the averages over generated data and the experimental values.

Stability estimators III: replicas vs data

Scatter correlation:

$$r\left[F^{(\mathrm{art})}\right] = \frac{\left\langle F^{(\mathrm{exp})} \left\langle F^{(\mathrm{art})} \right\rangle_{\mathrm{rep}} \right\rangle_{\mathrm{dat}} - \left\langle F^{(\mathrm{exp})} \right\rangle_{\mathrm{dat}} \left\langle \left\langle F^{(\mathrm{art})} \right\rangle_{\mathrm{rep}} \right\rangle_{\mathrm{dat}}}{\sigma_{s}^{(\mathrm{exp})} \sigma_{s}^{(\mathrm{art})}}$$

where the scatter variances are defined as

$$\begin{split} \sigma_s^{(\mathrm{exp})} &= \sqrt{\left\langle \left(F^{(\mathrm{exp})} \right)^2 \right\rangle_{\mathrm{dat}} - \left(\left\langle F^{(\mathrm{exp})} \right\rangle_{\mathrm{dat}} \right)^2} \;, \\ \sigma_s^{(\mathrm{art})} &= \sqrt{\left\langle \left(\left\langle F^{(\mathrm{art})} \right\rangle_{\mathrm{rep}} \right)^2 \right\rangle_{\mathrm{dat}} - \left(\left\langle \left\langle F^{(\mathrm{art})} \right\rangle_{\mathrm{rep}} \right\rangle_{\mathrm{dat}} \right)^2} \;. \end{split}$$

- $r\left[\sigma^{(\mathrm{art})}\right] r\left[\rho^{(\mathrm{art})}\right] r\left[\mathrm{cov}^{(\mathrm{art})}\right]$ are defined in the same way.
- The scatter correlation indicates the size of the spread of data around a straight line. Specifically $r\left[\sigma^{(\operatorname{art})}\right]=1$ implies that $\left\langle\sigma^{(\operatorname{art})}_i\right\rangle$ is proportional to $\sigma^{(\exp)}_i$.

- ullet Difficult to give a statistical measure of theoretical error: check that the final result depend within 2σ on theoretical assumptions.
- E.g. choice of the initial parametrisation:

$$d[q] = \sqrt{\left\langle rac{\left(q_i^{(1)} - q_i^{(2)}
ight)^2}{(\sigma_i^{(1)})^2 + (\sigma_i^{(2)})^2}
ight
angle_{ ext{dat}}},$$

- $q_i^{(1)}$, $q_i^{(2)}$ predictions for the *i*-th data point in the two fits, $\sigma_i^{(1)}$, $\sigma_i^{(2)}$ predictions for the corresponding statistical uncertainties.
- The results of the first and second fit are statistically equivalent if d[q] = 1 on average.
- The same must be done for the choice of kinematical cuts, random seeds, preprocessing exponents...

Neural Network and Training Algorithm

- Set neural network parameters randomly.
- Make clones of the parameter vector and mutate them.
- Evaluate the figure of merit for each clone:

$$\chi^{2(k)} = \sum_{i,j}^{N_{\text{dat}}} (F_i^{(\text{dat})(k)} - F_i^{(\text{net})(k)}) \operatorname{cov}_{ij}^{-1} (F_j^{(\text{dat})(k)} - F_j^{(\text{net})(k)})$$

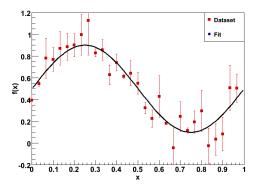
• Select the best ones and iterate the procedure until a stability is reached.

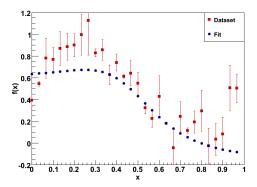
PROs

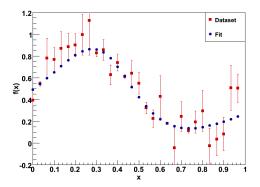
- The possibility of getting trapped in a local minimum is reduced.
- Allows to minimise the fully correlated χ^2 .

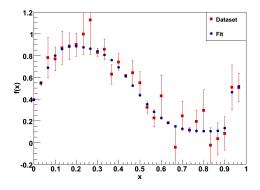
CONs

- It is monotonically decreasing by construction.
- It risks to converge slowly if the parameters ar not properly tuned.









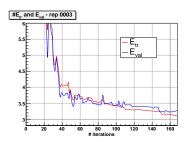
- Need a redundant parametrization to avoid excessive constraining
- Need a way of stopping the fit before overlearning sets in

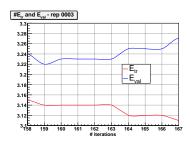


How to avoid Overlearning?

Stopping criterion based on Training-Validation separation

- * Divide data in two sets: training and validation.
- * Minimisation is performed only on the training set. The validation χ^2 for the set is computed.
- * When the training χ^2 still decreases while the validation χ^2 stops decreasing \to STOP.





The Evolution Code

- Observables are a convolution over x of PDFs and Coefficient Functions.
- Each observable is a particular linear combination of $(2n_f + 1)$ parton distributions.
- Data are given at various scales \rightarrow Solve DGLAP eqns and evolve from the initial parametrisation scale Q_0^2 to the experimental one.
- Theory: higher perturbative orders, resummations, higher twists, nuclear corrections, heavy quark threshold...

We want → Mellin space evolution.

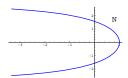
$$\tilde{\Gamma}(N,\alpha_s(Q^2),\alpha_s(Q_0^2)) = C(N,\alpha_s(Q^2)) \, \Gamma(N,\alpha_s(Q^2),\alpha_s(Q_0^2))$$

We do not want → Complex neural networks.

$$q(x,Q^2) = \int_x^1 \frac{dy}{y} \, \tilde{\Gamma}(y,\alpha_s(Q^2),\alpha_s(Q_0^2)) \, q(\frac{x}{y},Q_0^2)$$

$$\tilde{\Gamma}(y,\alpha_s(Q^2),\alpha_s(Q_0^2)) = \frac{1}{2\pi i} \int_C dN \, x^{-N} \, \tilde{\Gamma}(N,\alpha_s(Q^2),\alpha_s(Q_0^2))$$

Evolution



$$\begin{split} q(x,Q^2) &= \gamma \, q(x,Q_0^2) + \int_x^1 \frac{dy}{y} \, \Gamma(y,a_s,a_0) \, \left[q \left(\frac{x}{y},Q_0^2 \right) - y \, q(x,Q_0^2) \right] \\ \gamma &= \int_{c-i\infty}^{c+\infty} \frac{dN}{2\pi i} \frac{\Gamma(N)}{1-N} - \int_0^x dy \, \Gamma(y,a_s,a_0). \end{split}$$

×	$e_{\mathrm{rel}}(u_{V})$	$e_{\mathrm{rel}}(d_{v})$	$e_{\mathrm{rel}}(\Sigma)$	$e_{ m rel}(g)$
$1 \cdot 10^{-7}$	$7.6 \cdot 10^{-5}$	$4.5 \cdot 10^{-7}$	$1.3 \cdot 10^{-4}$	$1.2 \cdot 10^{-4}$
$1 \cdot 10^{-6}$	$8.3 \cdot 10^{-6}$	$3.2 \cdot 10^{-6}$	$1.2 \cdot 10^{-5}$	$1.7 \cdot 10^{-5}$
$1 \cdot 10^{-5}$	$4.7 \cdot 10^{-6}$	$1.4 \cdot 10^{-5}$	$1.5 \cdot 10^{-6}$	$2.2 \cdot 10^{-5}$
$1 \cdot 10^{-4}$	$3.3 \cdot 10^{-6}$	$3.3 \cdot 10^{-6}$	$1.4 \cdot 10^{-5}$	$4.8 \cdot 10^{-6}$
$1 \cdot 10^{-3}$	$1.3 \cdot 10^{-1}$	$9.7 \cdot 10^{-6}$	$2.6 \cdot 10^{-6}$	$1.5 \cdot 10^{-5}$
$1 \cdot 10^{-2}$	$2.9 \cdot 10^{-6}$	$1.6 \cdot 10^{-5}$	$5.5 \cdot 10^{-6}$	$4.9 \cdot 10^{-6}$
$1 \cdot 10^{-1}$	$7.9 \cdot 10^{-6}$	$1.1 \cdot 10^{-5}$	$5.2 \cdot 10^{-6}$	$3.8 \cdot 10^{-6}$
$3 \cdot 10^{-1}$	$1.4 \cdot 10^{-5}$	$2.7 \cdot 10^{-5}$	$1.8 \cdot 10^{-6}$	$3.6 \cdot 10^{-6}$
$5 \cdot 10^{-1}$	$2.8 \cdot 10^{-7}$	$1.0 \cdot 10^{-5}$	$1.7 \cdot 10^{-6}$	$6.4 \cdot 10^{-6}$
$7 \cdot 10^{-1}$	$9.0 \cdot 10^{-6}$	$7.3 \cdot 10^{-6}$	$8.7 \cdot 10^{-6}$	$7.6 \cdot 10^{-6}$
$9 \cdot 10^{-1}$	$1.1 \cdot 10^{-5}$	$6.0 \cdot 10^{-6}$	$1.0 \cdot 10^{-5}$	$7.9 \cdot 10^{-6}$

Table: LH benchmark vs NNPDF output for u_v , d_v , Σ and g distributions. **NLO** accuracy, **VFN** scheme, **truncated** solution. Inversion with FT algorithm.

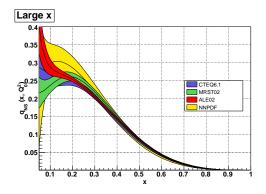
Non singlet fit

Determination of

$$T_3(x, Q_0^2) \equiv (u + \bar{u} - d - \bar{d})(x, Q_0^2)$$

at $Q_0^2 = 2 \text{GeV}^2$ at LO, NLO, NNLO.

 \bullet DATA SETS: $F_2^{\,p}(x,Q^2) - F_2^{\,d}(x,Q^2)$ BCDMS and NMC



See hep-ph/0701127 for all technical details

