# On determinations of the strong coupling constant from hadronic data

High Energy Physics Seminar, Cavendish Laboratory

Zahari Kassabov June 12th 2018

Cavendish Laboratory, University of Cambridge







## Introduction: Theoretical uncertainties at the LHC

The main uncertainties in a theoretical calculation of a process at the LHC:

- Missing higher orders uncertainties (MHOU).
- Uncertainties on the PDF.
- Uncertainties on  $\alpha_{\rm S}$ .

Example: ggH :



Figure 1: (S.Forte, Lattice 2017)

- Difficult to provide a statistical interpretation to all three of these uncertainties. Rely on a decision *by committee*.
- Improvements driven by:
  - Better calculations: Generally decrease uncertainty.
  - More reliable methodologies: Can decrease or increase uncertainty.
- In general, recent improvements in MHOU and PDF, not in  $\alpha_{\text{S}}$ .
- Reliable theory (and th uncertainties) crucial for LHC programme.

## Estimation of Missing Higher Order Uncertainties

MHOU: Usually estimated trough scale variations.

- Normative source: Yellow Report 4.
- Central scale setting can make a large difference [e.g jets at NNLO, Currie, et al arXiv:1704.00923].
- Size of the scale variation taken by convention (usually 2 and  $\frac{1}{2}$  for  $\mu_r$  and  $\mu_f$ ).
- Procedure to combine the two scale variations.
  - E.g.  $\Delta_{MHOU} = \left| \frac{\left[ \sigma(2\mu_f, 2\mu_r) \sigma(\frac{1}{2}\mu_f, \frac{1}{2}\mu_r) \right]}{2} \right| \text{ vs max} \text{min of 7-point variation.}$
  - $\cdot\,$  Uniform vs Gaussian distribution.
- Correlations in multi-process problems: Not well studied.

Improvements driven by calculating higher orders [e.g. ggH at N<sup>3</sup>LO, Anastasiou et al, arXiv:1602.00695].

### PDFs and PDF uncertainties

- Normative source: PDF4LHC Recommendation [arxiv:1510.03865].
- PDF uncertainties propagate experimental uncertainties to best fit of unbiased interpolants. MHOU not considered at all!
- Improvements driven by methodology, new data and theoretical calculations. Leads to agreement between groups.



Figure 2: Tie-Jiun Hou, DIS 2018 (Unpublished)

## Status of $\alpha_{\rm S}$

- Normative source: PDG combination [PDG review 2016, Bethke, Dissertori, Salam].
- World average on several determinations. Philosophy:

"...be as neutral as possible with regard to disputes in the community about different determinations, with uniform prescriptions applied to all reasonable determinations."

Salam, arxiv:1712.05165

• Current value

$$\alpha_{\rm S}(M_Z^2) = 0.1181 \pm 0.0013$$

• No updates in 2017 and 2018.

## **PDG** combination



- Only NNLO or better determinations considered.
- "Pre averaging": Take the unweighted mean and the mean error from each process.
- Final number obtained as a weighted " $\chi^2$  average" over the processes.

G.Altarelli, 2013

G.Altarelli, 2013

• Identifying most reliable determinations largely a sociological problem.

G.Altarelli, 2013

- Identifying most reliable determinations largely a sociological problem.
- PDG world average can be improved by:
  - Updates to individual determinations.
  - Identifying flawed determinations (potentially bigger effect).

G.Altarelli, 2013

- Identifying most reliable determinations largely a sociological problem.
- PDG world average can be improved by:
  - Updates to individual determinations.
  - Identifying flawed determinations (potentially bigger effect).
- How reliable is the determination of  $\alpha_{S}$  from PDFs?

- Basic concept: Find best fit in the combined space of ( $\alpha_s$ , PDF) of a large body of experimental data.
- Compared to Electroweak precision fit [Blas et al, arXiv:1608.01509], the larger dataset implies:
  - More challenging to achieve a good description of all processes (e.g. low scale DIS affected by both MHOU and higher twists).
  - Theoretical problems likely to average out to some extend [Carrazza, Forte, ZK, Rojo, Rottoli, arxiv:1803.07977].
  - More data and more dependence on  $\alpha_{\rm S}$  implies more precision.

#### Challenges extracting $\alpha_{\rm S}$ from PDFs

- PDF parametrization may bias the  $\alpha_{\rm S}$  value.
- Correct treatment of experimental systematics (particularly normalization uncertainties).
- Hidden uncertainties in theoretical description of PDFs (e.g. heavy quark treatment).
- Inclusion of PDF uncertainty in the  $\alpha_{\rm S}$  determination.
- Accurate estimation of missing higher order uncertainties

### Challenges extracting $\alpha_{\rm S}$ from PDFs

- + PDF parametrization may bias the  $\alpha_{\rm S}$  value. [NNPDF]
- Correct treatment of experimental systematics (particularly normalization uncertainties). [NNPDF]
- Hidden uncertainties in theoretical description of PDFs (e.g. heavy quark treatment). [NNPDF 3.1]
- Inclusion of PDF uncertainty in the  $\alpha_{\rm S}$  determination. [NNPDF 3.1- $\alpha_{\rm S}$ ]
- Accurate estimation of missing higher order uncertainties [TO DO]

Made notable progress since latest NNPDF determination in 2011.

- +  $\alpha_{\rm S}$  determination based on NNPDF3.1
- Can collider determinations of  $\alpha_{\rm S}$  be really independent on those from PDFs?
- Can we *measure* theoretical uncertainties?

Precision determination of the strong coupling constant within a global PDF analysis

NNPDF Collaboration, arxiv:1802.03398

- 3979 data points.
- Including simultaneously differential top, *Z pT* and inclusive jet data for the first time.
- Exact NNLO theory
- Method to effectively fit  $\alpha_{\rm S}$  and the PDF simultaneously (PDFs effectively nuisance parameters).

Precision determination of the strong coupling constant within a global <del>PDF QCD</del> analysis

NNPDF Collaboration, arxiv:1802.03398

- 3979 data points.
- Including simultaneously differential top, *Z pT* and inclusive jet data for the first time.
- Exact NNLO theory
- Method to effectively fit  $\alpha_{\rm S}$  and the PDF simultaneously (PDFs effectively nuisance parameters).

#### Dataset in NNPDF 3.1



 $\alpha_{\rm S}$  dependence though:

- PDF evolution (large scale differences advantageous).
- Direct dependence (e.g. *tī* data)
- Higher order QCD corrections.

Our data has experimental uncertainties.

- We view the data as random variables from the distribution  $\mathcal{N}(d_i, \Sigma_{i,j})$ 
  - $d_i$  is the experimentally measured central value for the point *i*.
  - $\Sigma_{ij}$  *a* covariance between the points *i*, *j*.
- We sample  $N_{\text{rep}}$  datasets (Monte Carlo pseudodata) from the distribution, and train neural networks "replicas"s to each dataset to minimize an error function  $\chi^2$ .
- PDF dependent quantities are calculated from statistics over the ensemble of replicas. E.g. "*PDF uncertainty*" is usually the standard deviation over the replicas.

## Simultaneous minimization of ( $\alpha_{\rm S}$ , PDF)

- Ideally would minimize simultaneously  $\alpha_{\rm s}$  and the PDF parameters in each replica.
  - Can't do easily because we use precomputed tables that depend on  $\alpha_{\rm s}.$

## Simultaneous minimization of ( $\alpha_{s}$ , PDF)

- Ideally would minimize simultaneously  $\alpha_{\rm s}$  and the PDF parameters in each replica.
  - Can't do easily because we use precomputed tables that depend on  $\alpha_{\rm s}.$
  - Solution: Repeat the fit for discrete values of  $\alpha_S$  to the same data replica  $\longrightarrow$  c-replica.

## Simultaneous minimization of ( $\alpha_{S}$ , PDF)

- Ideally would minimize simultaneously  $\alpha_{\rm s}$  and the PDF parameters in each replica.
  - Can't do easily because we use precomputed tables that depend on  $\alpha_{\rm s}.$
  - Solution: Repeat the fit for discrete values of  $\alpha_{\rm S}$  to the same data replica  $\longrightarrow$  c-replica.
- Each c-replica has a  $\chi^2(\alpha_S)$  profile.



## Simultaneous minimization of ( $\alpha_{S}$ , PDF)

- Ideally would minimize simultaneously  $\alpha_{\rm s}$  and the PDF parameters in each replica.
  - Can't do easily because we use precomputed tables that depend on  $\alpha_{\rm s}.$
  - Solution: Repeat the fit for discrete values of  $\alpha_S$  to the same data replica  $\longrightarrow$  c-replica.
- Each c-replica has a  $\chi^2(\alpha_S)$  profile.
- Each minimum yields one sampled  $\alpha_{\rm S}$  value.



#### We obtain

$$\alpha_{\rm S}^{\rm NNLO}(M_Z) = 0.11845 \pm 0.00052^{(\rm exp)}(0.4\%)$$

- Experimental uncertainty is the standard deviation over the ensemble of c-replicas of the  $\arg \min_{\alpha_S} \chi^2(\alpha_S)$ .
- Very small uncertainty (compared to other determinations in the PDG). But need to assess:
  - Methodological systematics.
  - Theory uncertainties.

#### Improvement by batch minimization

- The NNPDF methodology discards replicas that do not meet certain convergence criteria (25-30%).
- $\chi^2(\alpha_S)$  noisy for each c-replica. Fit quality depends on:
  - Initial conditions.
  - Cross validation split
- Solution: Repeat each fit three times, and take the minimum (requiring at least 2 successful replicas).



## Effect of batch minimization

- Small on central value  $\sim$  O(0.4)  $imes \Delta^{(exp)}$ .
- · Significative on uncertainty. Reduction to up to 27%.
- · Convergence already good with two batches.



We estimate the uncertainties due to fitting a finite number of replicas by bootstrapping.

- 1. Take the set of *N* minima.
  - 1.1 Sample with replacement from it *M* sets of *N* values, with *M* large.
  - 1.2 Compute the *M* means of each of the *M* sets.
  - 1.3 Compute the standard deviation of the *M* means.

$$\Delta_{stat} = 3 \times 10^{-5}$$

Effect negligible compared to the PDF uncertainty.

- Select c-replicas that contain at least *N*<sub>min</sub> successfully fitted values.
- Choose  $N_{\min}$  to minimize the bootstrapping uncertainty.
  - $\cdot\,$  Tradeoff between statistics and quality of the parabolas.
- Result stable for reasonable values  $N_{\min}$ .

$N_{\min}$	$\alpha_s \left( m_Z \right)$	$N_{\rm rep}$	$\Delta_{\alpha_s}$
18	$0.11842 \pm 0.00031 \ (0.3\%)$	12	0.00009
15	$0.11844 \pm 0.00044 ~(0.4\%)$	92	0.00005
6	$0.11845 \pm 0.00052~(0.5\%)$	379	0.00003
3	$0.11844 \pm 0.00056 ~(0.5\%)$	400	0.00003

## Treatment of normalization uncertainties

- It is well known [D'Agostini, 2003] that minimizing the experimental  $\chi^2$  is biased in the presence of normalization uncertainties.
  - Maximum likelihood estimator is not an unbiased estimator.
- NNPDF uses the  $t_0$  procedure [arxiv:0912.2276] (fix normalization from the result of a previous fit).
- $\alpha_{\rm S}$  fit heavily biased when the experimental definition of  $\chi^2$  is used inconsistently to minimize  $\alpha_{\rm S}$ .



- Checked many other possible sources of systematics.
  - Assumption that  $\chi^2(\alpha_s)$  is parabolic.
  - Dependence on the extreme values.
  - Myriad of variations of replica selection.
  - Effect of  $t_0$  procedure.
- Overall all effects much smaller than experimental uncertainty. Methodological effects conservatively estimated at  $\sim 10^{-4}(0.09\%).$

## Theoretical uncertainties that are under control

- Theoretical effects necessarily have to affect the PDFs in order to affect our determination.
- But not sufficient: Dependence on  $\alpha_{\rm S}(M_Z)$  required.
- Many have been studied and found to be smaller than experimental uncertainties:
- **Higher twist** found to be small compared to experimental uncertainties [arxiv:1303.1189].
- **Charm mass** greatly improved by parameterizing the charm PDF [arXiv:1605.06515].
- **Electroweak corrections** Kept under control with suitable cuts [arxiv:1706.00428].
- **Nuclear corrections** Studied in [arxiv:1706.00428] and found to be small compared to experimental uncertainties.

In conclusion MHOU highly likely to be the dominant theoretical uncertainty.

No reliable method known for PDF based quantities. We have:

 $\alpha_{\rm S}^{\rm NNLO}(M_Z) = 0.1184 \pm 5 \times 10^{-4}$ 

$$\alpha_{\rm S}^{\rm NLO}(M_Z) = 0.1206 \pm 6 \times 10^{-4}$$

- PDF fits with scale variations currently in early stage.
- Methods based on continuation of perturbative series such as Cacciari-Houdeau [arXiv:1105.5152] hampered by:
  - Poor fit quality at LO and even at NLO.
  - Lack of unique series expansion (many processes involved).
- $\cdot\,$  CH yields a theoretical uncertainty of 4  $\times\,10^{-4},$  smaller that the experimental ones. Likely too optimistic.

For lack of better options, make it be crude and conservative.

$$\Delta \alpha_{\mathrm{S}}^{\mathrm{th}} = \frac{1}{2} \left| \alpha_{\mathrm{S}}^{\mathrm{NLO}} - \alpha_{\mathrm{S}}^{\mathrm{NNLO}} \right| = 0.0011 (0.9\%)$$

• Uncertainty likely overblown by the poor fit quality at NLO.

	$\min_{lpha_{\sf S}} \chi^2(lpha_{\sf S}^{({\sf central})})/N_{\sf data}$
NLO	5014/3979=1.26
NNLO	4814/3979=1.21

 $\cdot$  Two times bigger than the experimental uncertainty.

## There must be a better way! A data driven way?

Let us make the following assumptions:

- Data entering the fit can be separated in *processes* where MHOU are largely uncorrelated.
  - In practice correlations expected because of the PDF.
  - Also must see how to deal with experimental correlations.
- One can define a preferred value of  $\alpha_{\rm S}$  for each process somehow.
- There are sufficient number of processes to make statistics.
- Differences in preferred values above experimental uncertainties are due to theoretical differences.

It follows that theoretical uncertainties can be estimated trough the dispersion of the preferred values.

• Strong assumptions, but is it really worse than scale variations?

## How NOT to define preferred values: The partial $\chi^2$

The global  $\chi^2$  minimizes to fit ( $\alpha_{\rm S}, {\sf PDF}$ ) looks like

$$\chi^{2}\left[\{\theta\}, \alpha_{\mathrm{S}}, \mathcal{D}\right] = \sum_{l,l=1}^{N_{\mathcal{D}}} \left(T_{l}\left[\{\theta\}, \alpha_{\mathrm{S}}\right] - D_{l}\right) C_{lj}^{-1}\left(T_{j}\left[\{\theta\}, \alpha_{\mathrm{S}}\right] - D_{j}\right)$$

where

- $\{\theta\}$  are the parameters of the PDF.
- +  $\mathcal D$  is the set of  $\mathit{N_{\mathcal D}}$  data points entering the fit.
- $T_I$  is the theoretical prediction for the data point indexed by I.
- $D_1$  is the experimentally measured value.
- C<sub>IJ</sub> measures the experimental covariance.

The profile of a c-replica is

$$\chi^{2}(\alpha_{s}) = \min_{\{\theta\}} \chi^{2}\left[\{\theta\}, \alpha_{s}, \mathcal{D}\right]$$

The partial  $\chi^2$  is:

$$\chi_{p}^{2}\left[\{\theta\}, \alpha_{S}, \mathcal{P}\right] = \sum_{l,l=1}^{N_{\mathcal{P}}} \left(T_{l}\left[\{\theta\}, \alpha_{S}\right] - D_{l}\right) C_{lj}^{-1}\left(T_{j}\left[\{\theta\}, \alpha_{S}\right] - D_{j}\right)$$

where we have replaced  $\mathcal{D}$  with  $\mathcal{P}$ , a subset of  $\mathcal{D}$  with  $N_{\mathcal{P}}$  points, and neglected the correlations between the points in  $\mathcal{P}$  and  $\mathcal{D}$ .

 $\cdot\,$  Up to missing correlations, for a set of processes that cover all  ${\cal D}$ 

$$\chi^2(\alpha_s) = \sum_p \chi^2_p(\alpha_s)$$

#### Results from partial $\chi^2$



• Define

$$\alpha_{\rm S}^{\rm p} = \arg\min_{\alpha_{\rm S}} \chi_{\rm p}^2(\alpha_{\rm S})$$

- Useful to estimate pulls qualitatively. E.g. can say that the LHC data likely contributes to increase  $\alpha_s(M_Z)$ .
- But quantitative calculations are seriously off.

By "collider determinations" I mean determinations of  $\alpha_s$  based on cross sections measured at hadron-hadron and hadron-lepton colliders that are used to constrain the strong coupling independently of a PDF fit.

G.Salam [arxiv:1712.05165]

However all "collider determinations" in practice amount to the partial  $\chi^2$  minimization above. Collider determinations depend on the PDF in two ways:

- Best fit PDF changes strongly with  $\alpha_{s}$ , i.e.  $\{\theta\}(\alpha_{s}) = \arg\min_{\{\theta\}} \chi^{2}(\alpha_{s}, \{\theta\}, \mathcal{D}).$
- Global fit quality  $\chi^2(\alpha_S)$  changes strongly with  $\alpha_S$ .

- Notable example CMS measurement of  $t\bar{t}$  [arXiv:1307.1907] computed at NNLO [Czakon, Fiedler, Mitov, arXiv:1303.6254] included in the PDG average as the only item in the collider determination category.
- Many others have appeared recently. For example, a determination using jets in DIS at NNLO [H1 Collaboration arxiv:1709.07251] at. The result, using NNPDF3.1 sets, is  $\alpha_{\rm S} = 0.1157 \pm 0.002^{\rm exp} \pm 0.003^{\rm th}$ . The central value is discarded by the NNPDF3.1 determination at  $5\sigma$ .

## Problems with the partial $\chi^2$

• Neglecting correlations is a bad approximation.

$$lpha_{
m S}^{({
m global})} - rg\min_{lpha_{
m S}} \sum_{
ho} \chi_{
ho}^2(lpha_{
m S}) \sim {
m std}_{
ho} lpha_{
m S}^{
ho}$$

Deviation of the order of the quantity we wanted to estimate.

- By construction it doesn't take into account the global fit quality [Z.K, arxiv:1802.05236].
- Preferred values extracted this way amount to a logical contradiction:
  - Choose a point in the phase space  $(\alpha_{s}^{p}, \arg\min_{\alpha} \chi^{2}(\{\theta\}, \alpha_{s}^{p}, \mathcal{D}))$ ,

best fit PDF at  $\alpha_s^p$ 

- $\cdot$  discarded by the data ( $\mathcal{D}$ ) ,
- used to construct constrain the PDF parameters  $\{\theta\}$ ,
- on which  $\alpha_{\rm s}^{\rm p}$  relies in the first place.

## Fit results outside the global best fit

- $\chi^2_p$  depends strongly on the rest of the data entering the fit.
- $\chi_p^2$  inversely correlated to relative weight of  $\mathcal{P}$  (given by the number of points). The more weight  $\longrightarrow$  the more advantageous it is to optimize for it at the expense of some other data  $\longrightarrow$  the smaller  $\chi_p^2$ .
- For example, note  $t\bar{t}$ . 26 points (0.6% of the total) are only described simultaneously with the rest of the data in a small range of  $\alpha_s$  values.



Can it be that for processes like  $t\bar{t}$ ,  $\chi_p(\alpha_S)$  changes much more quickly than the global  $\chi^2(\alpha_S)$ . Made an experiment to test it:

- The further away from the global minimum, the larger the space of PDs with the same  $\chi^2(\alpha_s)$ .
- Can alter the fit in a way that changes substantially  $\chi^2_{\rm p}$  and very little the global  $\chi^2.$
- Set up:
  - Take the 26 points of  $t\bar{t}$  data and copy them 15 times in a fit at  $\alpha_{\rm S}(M_Z) = 0.121$ . The target function becomes  $\chi'^2 = \chi^2 + 14\chi_{t\bar{t}}^2$ .
  - Now *t*{t} has more weight in the fit.

$\chi^2/d.o.f.$	$\alpha_{\rm s}=0.118$	$\alpha_{ m s}=$ 0.121 default	$\alpha_{\rm s} = 0.121$ weighted $\overline{t}t$
Total	1.162	1.212	1.228
tī	1.07	1.42	1.02

- Very small change in the total  $\chi^2$ .
- Can bend the PDF to describe the top data perfectly at a larger value of  $\alpha_{\rm S}$ .

## Changes in the PDFs

- Bigger  $\alpha_{\rm S}$  compensate with smaller gluon PDF in the relevant kinematic region.



- Still compatible within uncertainties.
- Rest of PDFs largely unchanged.
- Conclusion: *t*t determination not independent on PDFs.

- It turns out that in our analysis the minimum for  $\chi^2_{t\bar{t}}$ ( $\alpha_S(M_Z) = 0.1183$ ) is very close to the global minimum  $\chi^2(\alpha_S(M_Z) = 0.11845)$ .
- Selecting a different process, we can engineer a fit where both  $\chi^2_p$  and  $\chi^2$  are lower than for the minimum of the partial  $\chi^2$ .
- For Z pT we have  $\alpha_{\rm S}^{Z \, pT} \approx 0.124$  in the default fit.
- Assign the 120 points of *Z pT* data a weight of 32 (this makes it weight as much as the rest of the data).

$\chi^2/d.o.f.$	$\alpha_{\rm s} = 0.120$ weighted Z pT	$\alpha_{\rm s}=$ 0.124 default
Total	1.226	1.281
ZpT	0.94	1.11

The weighted fit agrees better both with the whole ensemble  $\mathcal{D}$  and the ZpT data. It is therefore a better value of  $\alpha_S$  from ZpT, no matter how you look at it.

• I consider this a proof that the partial  $\chi^2$  minimization entails a logical contradiction, since in principle it is possible to find a better fit.

- We could exploit similar weighting strategy to define dataset dispersion [Z.K. arxiv:1802.05236].
- Associate a large enough weight to each dataset to force  $\chi_p^2(\alpha_S)/d.o.f. \simeq 1$  (statistical minimum when cross validation is applied) in a large range of  $\alpha_S$ .
- Define preferred value as the one where the rest of the data agrees the best, i.e.  $\alpha_{\rm S}^{({\rm preferred } p)} = \min_{\alpha_{\rm S}} \chi^2(\alpha_{\rm S})^{({\rm weighted } p)}$ .

#### Advantages of definition of preferred value

- Ideally defined as best fit to all the data, restricted to describing perfectly the specific dataset (weight is an implementation detail).
- $\cdot\,$  Explicitly depends on all the data in the problem  $\mathcal{D}.$
- Dispersion of preferred values over datasets (above experimental uncertainties) can be interpreted as mainly coming from theory.
- Possible formula

$$\Delta^{th} = \Delta^{exp} \frac{1}{N_p - 1} \sqrt{\sum_{p} \frac{\left(\alpha_s - \alpha_s^{(\text{preferred } p)}\right)^2}{\frac{1}{2}(\Delta^{2,exp} + \Delta^{2(exp,\text{preferred},p)})}}$$

- Determinations of  $\alpha_{\rm S}$  from global QCD fits (i.e. PDFs) have interesting characteristics to constrain its value.
- "Collider determinations" are flawed, and not independent on PDFs. Should not be used for World averages.
- It is interesting to explore data driven methods to study theory uncertainties, e.g. the dataset dispersion outlined here.

## Thank you!